

# Length-weight analysis for southern hake stock

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2022-03-09

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Nowaday, the length-weight relationship carried out in 1999 is used. This relationship provides global (not year specific) estimates of the  $a$  and  $b$  parameters ( $W_i = aL_i^b$ ,  $i = 1, \dots, N$ , being  $N$  the total sample size). More precisely, the actual values are  $a = 0.00659$  and  $b = 3.01721$ . Hence, in the current study a review of length-weight relationship is done addressing the following tasks:

- Estimation of global (not year specific)  $a$  and  $b$  using the updated length-weight data base.
- Estimation of year specific  $b$  (and common  $a$ ).
- Estimation of global  $a$  and  $b$  by sex.
- Analysis of previous results and derived proposal.

**Models:** The typical length-weight model is  $W_i = aL_i^b$ , where  $a$  and  $b$  are parameters to be estimated. If we take logarithms (with base 10) in both sides, we obtain  $\log_{10}(W_i) = \log_{10}(a) + b \log_{10}(L_i)$ . This model can be fitted in R using the common functions for linear models:

$$\log_{10}(W_i) = \log_{10}(a) + b \log_{10}(L_i) + \epsilon_i,$$

where  $\epsilon_i$  normally distributed with mean 0 and variance  $\sigma^2$ .

## Prepare data

We read the data files: Portugal and Spain data. We have then three data sets, **dataS** corresponds to the Spain data, **dataP** corresponds to Portugal data, and finally **data** contains both data sets.

```
##   year month   tl   tw gw sex source      area prof M   Y   logL   logW
## 1 2009     7 36.7 385 NA   F market 9a-S_cadiz  NA 7 2009 1.564666 2.585461
## 2 2009     7 29.0 210 NA   F market 9a-S_cadiz  NA 7 2009 1.462398 2.322219
## 3 2009     7 36.0 330 NA   F market 9a-S_cadiz  NA 7 2009 1.556303 2.518514
## 4 2009     7 35.2 414 NA   F market 9a-S_cadiz  NA 7 2009 1.546543 2.617000
## 5 2009     7 30.9 245 NA   F market 9a-S_cadiz  NA 7 2009 1.489958 2.389166
## 6 2009     7 35.6 345 NA   F market 9a-S_cadiz  NA 7 2009 1.551450 2.537819
```

We unify the *area* variable defining the categories northwestern Cantabrian fishing grounds (Spain data, termed “CNO”) and Portugal.

```
##
##      CNO Portugal
##      41276    48331
```

The time series is reduced starting in 1982. Note that data for 2020 is not considered since not enough sampling has been carried out in such year.

```

##   Min. 1st Qu. Median   Mean 3rd Qu.   Max.
##   1982    2004    2009    2007    2013    2019

```

## Exploratory

### General

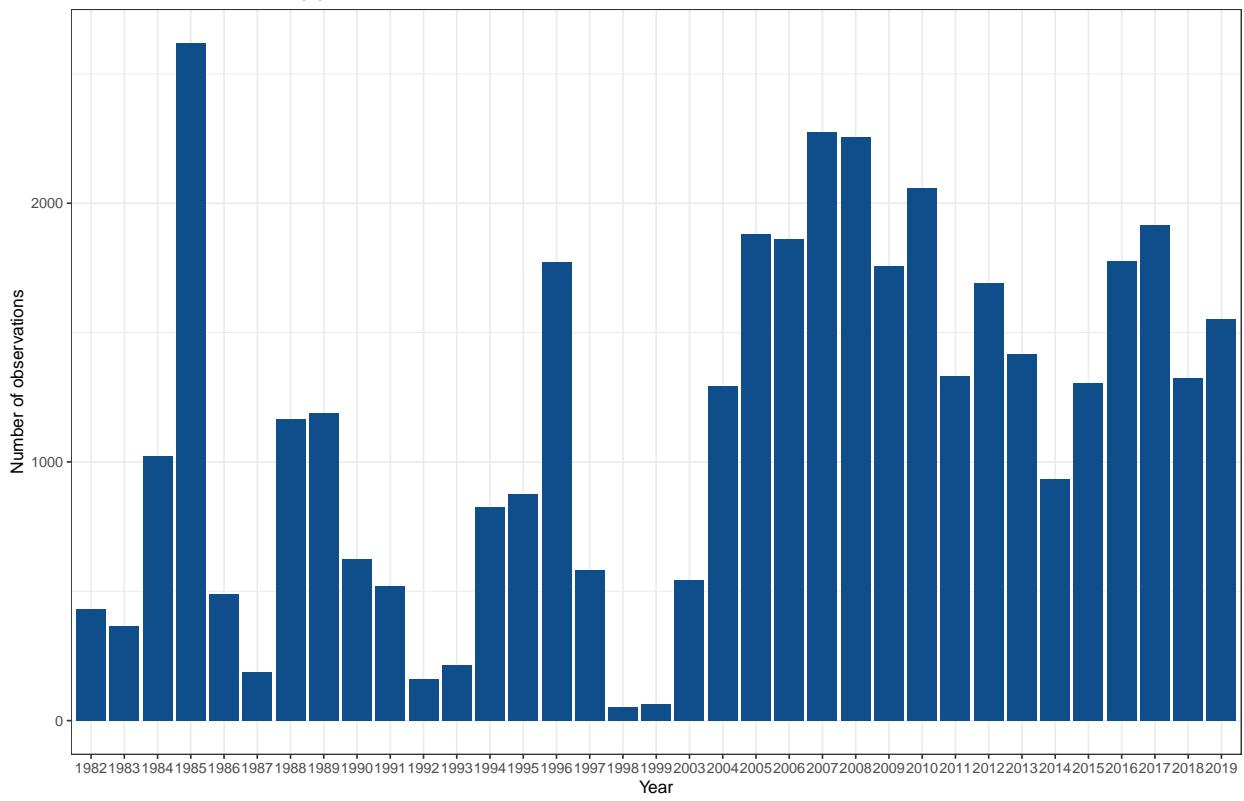
We check the number of individuals by area. Portugal data cover from 1996 to 2018, whereas Spain data cover from 1982 to 2019. Note that in 2000, 2001, 2002 there are not individuals in CNO (Cantabrian fishing grounds), also no data is available for Portugal for 1997-1999 and 2019.

```

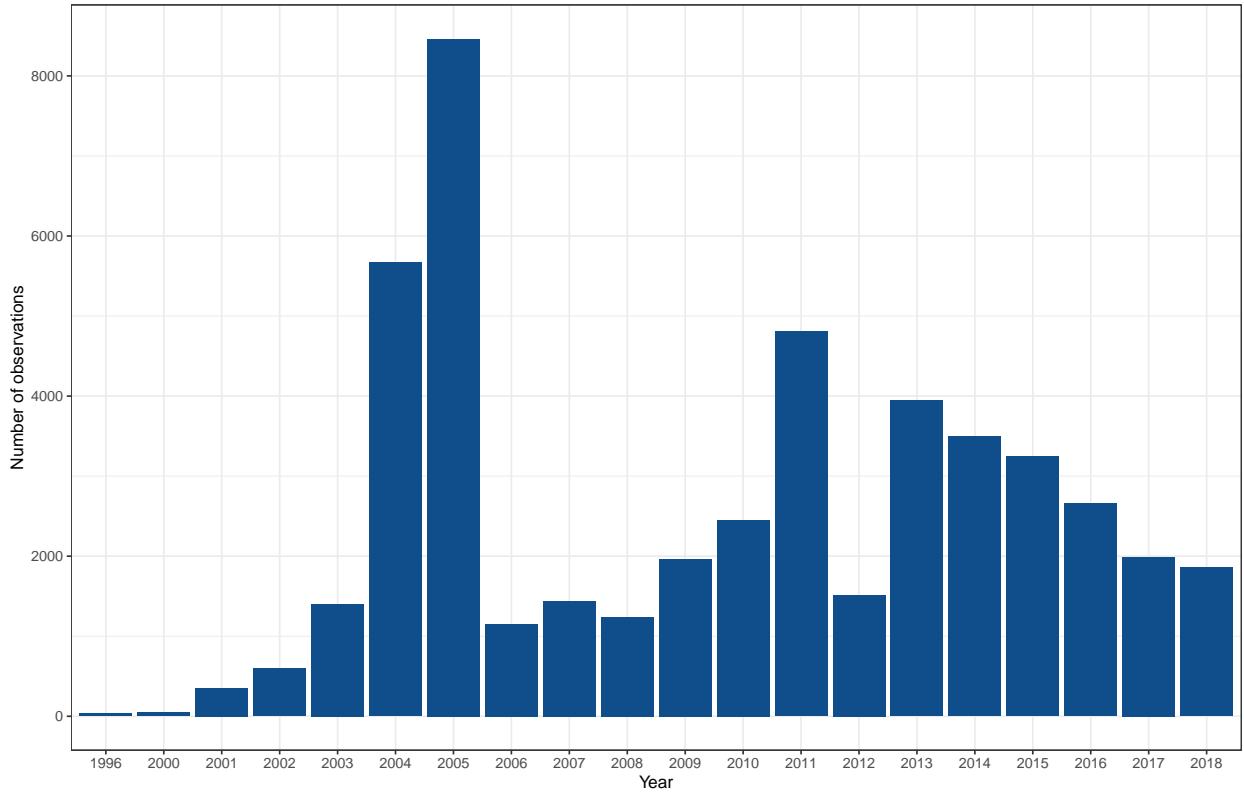
##
##          CNO Portugal
##  1982    431      0
##  1983    364      0
##  1984   1023      0
##  1985   2619      0
##  1986    490      0
##  1987    187      0
##  1988   1165      0
##  1989   1187      0
##  1990    623      0
##  1991    519      0
##  1992    160      0
##  1993    214      0
##  1994    827      0
##  1995    874      0
##  1996   1773     35
##  1997    581      0
##  1998    53       0
##  1999    63       0
##  2000     0      47
##  2001     0     350
##  2002     0     596
##  2003   542    1401
##  2004  1295    5670
##  2005  1880    8464
##  2006  1862   1150
##  2007  2275   1443
##  2008  2254   1243
##  2009  1757   1963
##  2010  2058   2455
##  2011  1333   4807
##  2012  1693   1516
##  2013  1418   3947
##  2014  934    3495
##  2015 1306    3243
##  2016 1777   2658
##  2017 1914   1991
##  2018 1323   1857
##  2019 1552      0

```

Number of observations by year for CNO

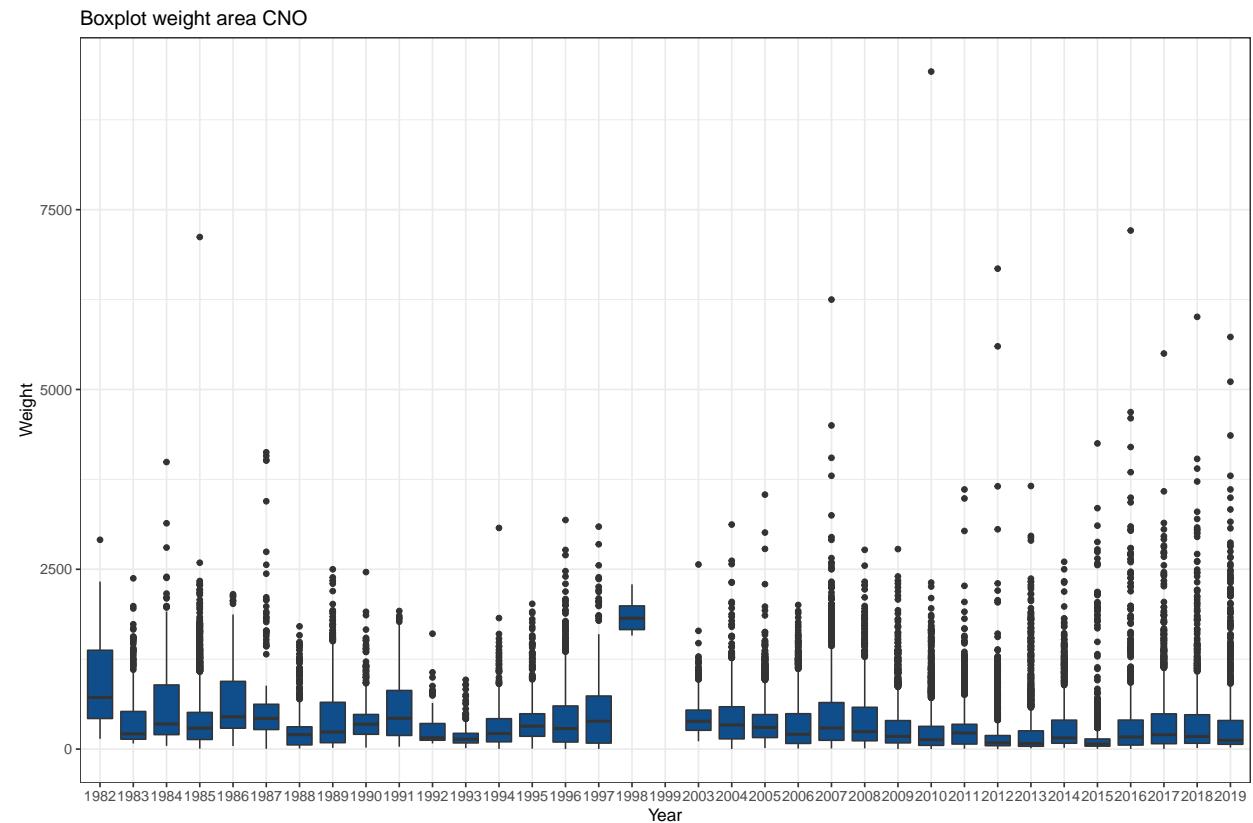
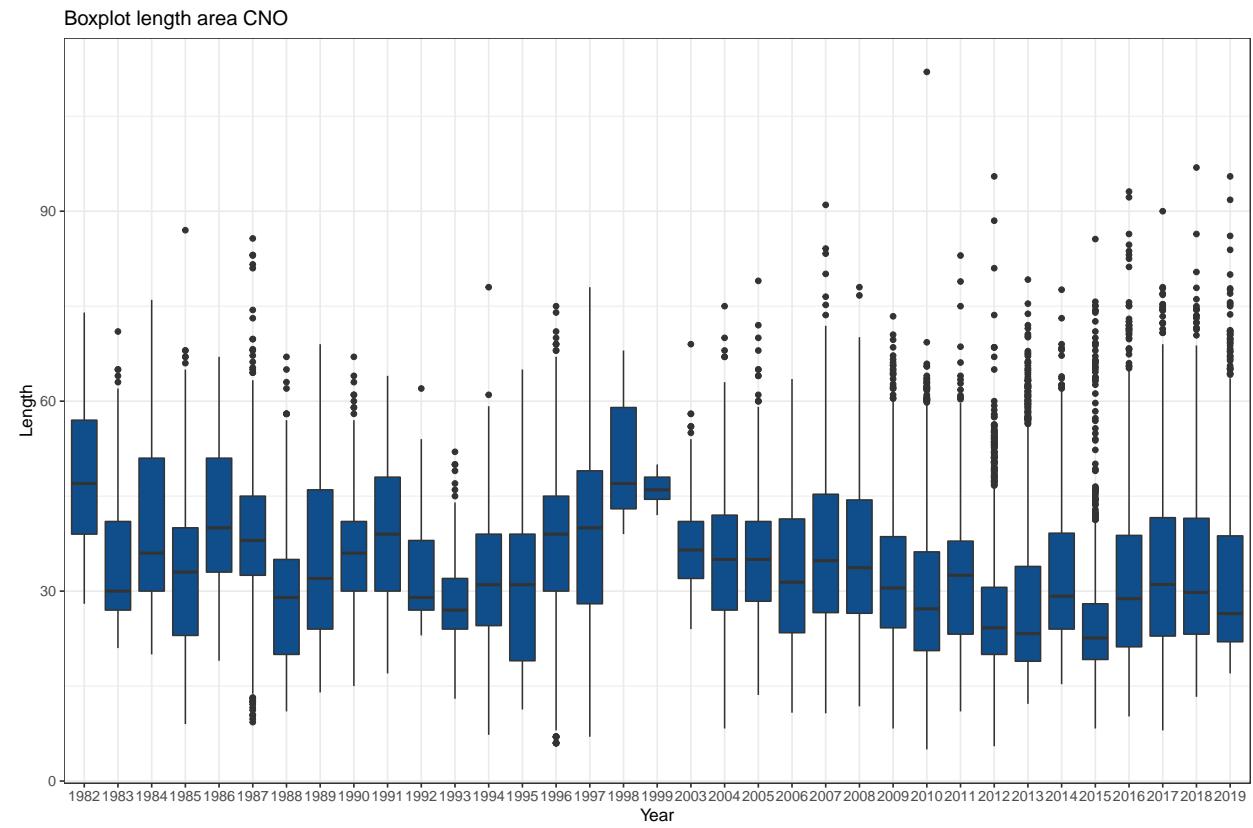


Number of observations by year for Portugal



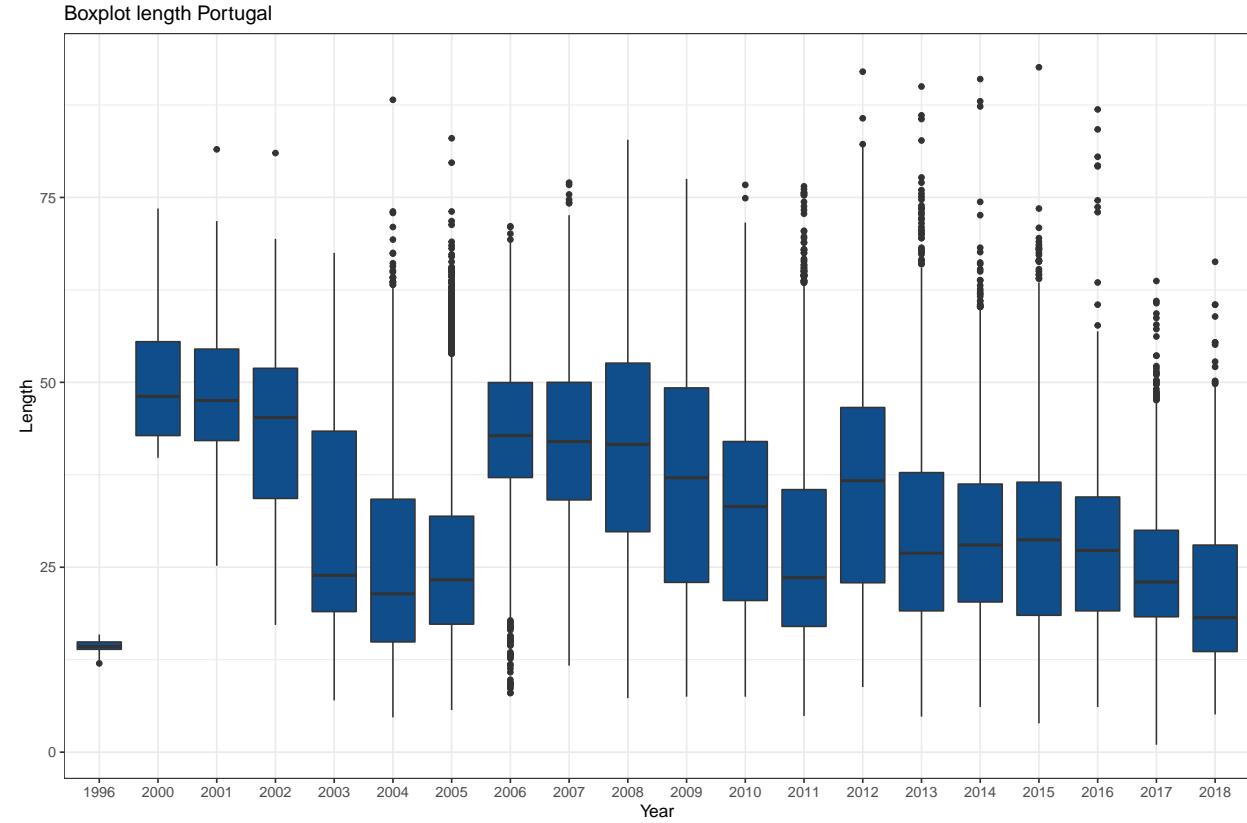
We look at length and weight along time (by grouped area). Two outliers in weight have been detected in

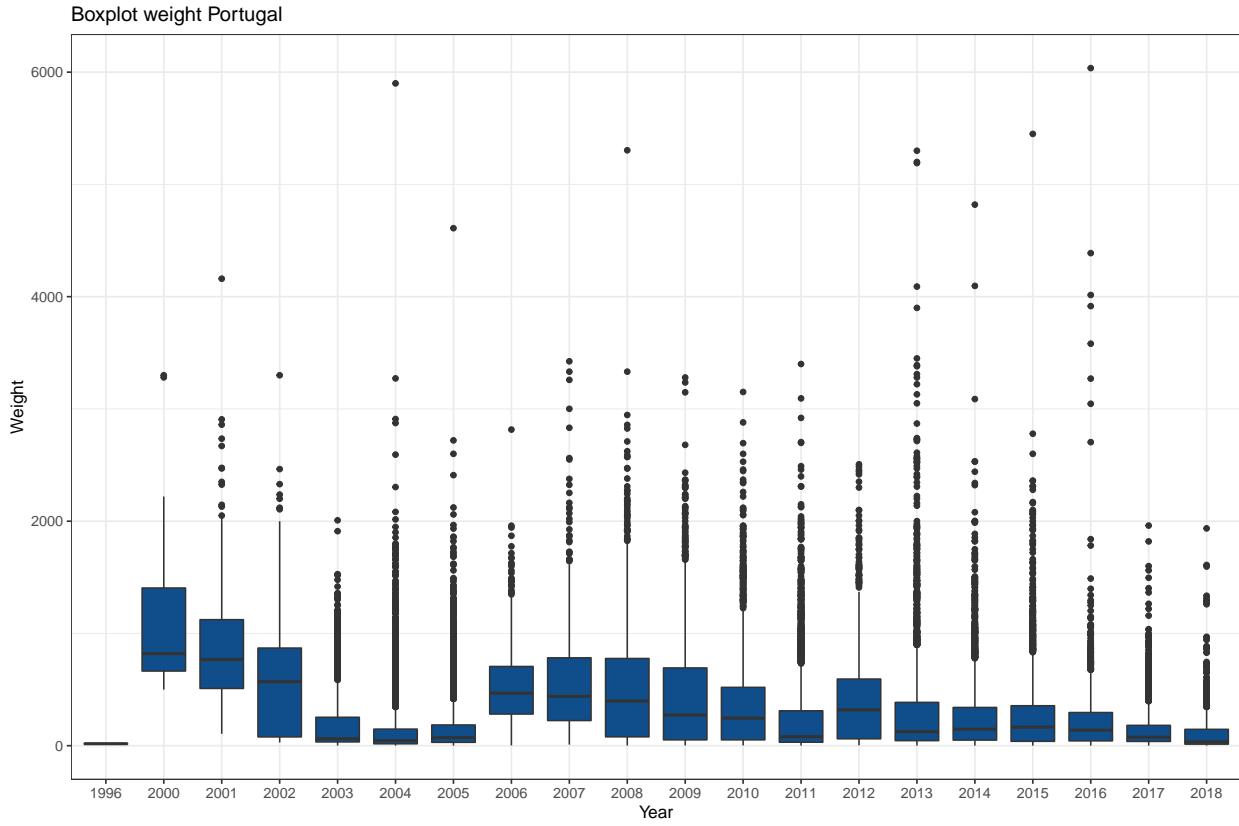
2010 (CNO area), and hence they have been removed (their values were 110211, 116611).



Year 1999 does not contain the weights for the corresponding lengths. Hence, 1999 for CNO is removed from the data base. Year 1998 has a particular behaviour and a small sample size (53 individuals).

```
##      year month tl tw gw sex source area prof M     Y      logL logW areaU
## 14363 1999      2 48 NA 648   I market  8c  NA 2 1999 1.681241   NA  CNO
## 14364 1999      2 49 NA 714   I market  8c  NA 2 1999 1.690196   NA  CNO
## 14365 1999      2 48 NA 686   I market  8c  NA 2 1999 1.681241   NA  CNO
## 14366 1999      2 50 NA 750   I market  8c  NA 2 1999 1.698970   NA  CNO
## 14367 1999      2 47 NA 693   I market  8c  NA 2 1999 1.672098   NA  CNO
## 14368 1999      2 46 NA 605   F market  8c  NA 2 1999 1.662758   NA  CNO
```

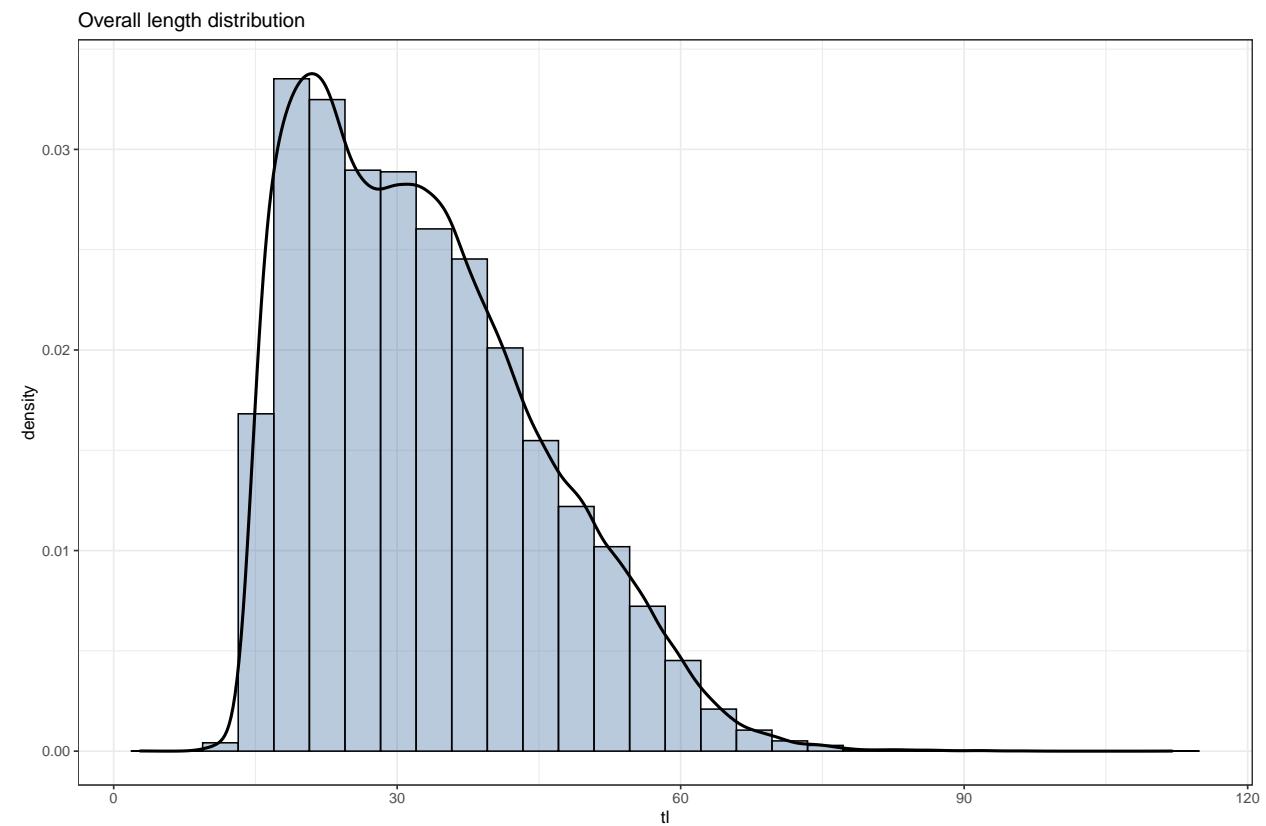


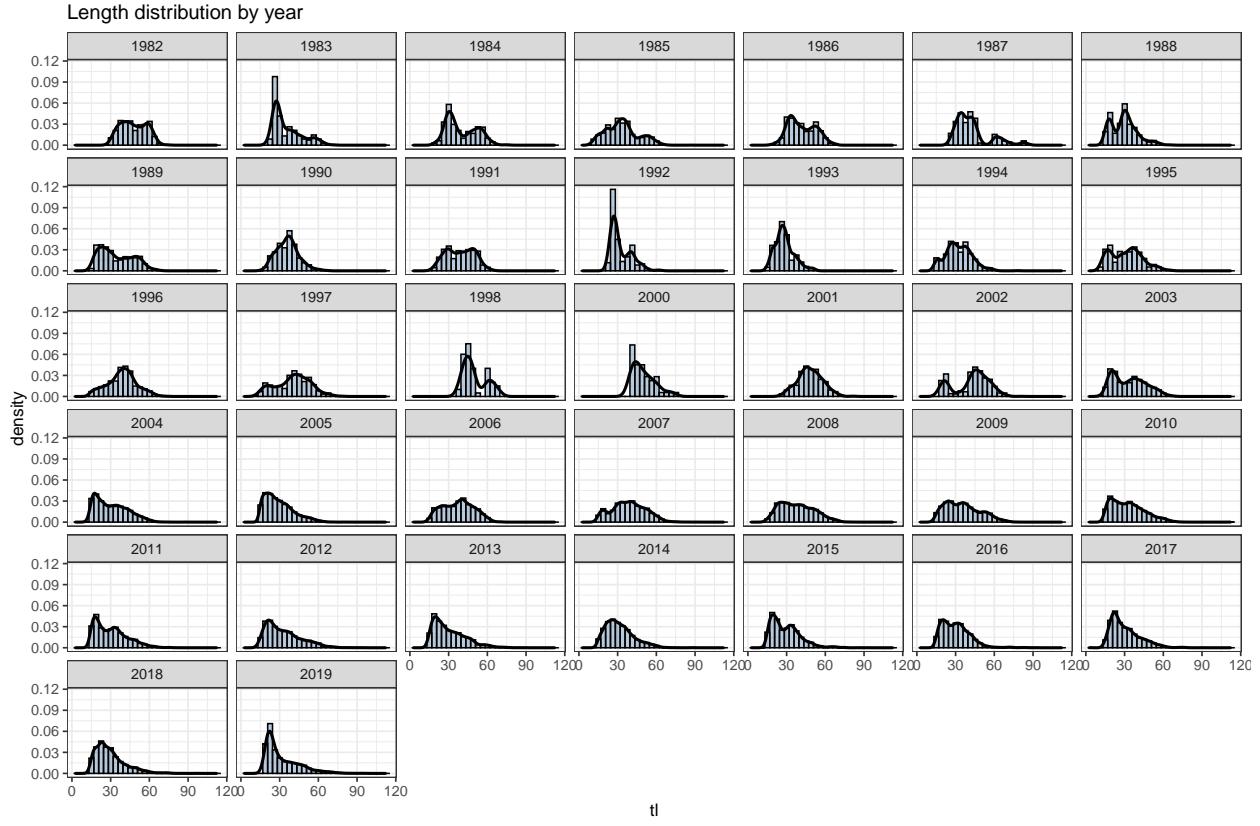


Year 1996 has a clear strange behaviour, then we eliminate the data of this year whose sample size is 35 (too small).

Finally, we decide to eliminate all individuals whose weight is less or equal than 18 gr since the measures for weights smaller than these ones are not trustworthy.

## Length distributions





## Length-weight model

### Global model

Estimation of global (not year specific)  $a$  and  $b$  using the updated length-weight data base.

```
fit.0 <- lm(logW ~ logL, data = data)
summary(fit.0)

##
## Call:
## lm(formula = logW ~ logL, data = data)
##
## Residuals:
##      Min       1Q   Median       3Q      Max 
## -0.90013 -0.03208 -0.00432  0.02778  2.99730 
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)    
## (Intercept) -2.398095   0.001881  -1275   <2e-16 ***
## logL         3.153611   0.001268   2488   <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.05646 on 76590 degrees of freedom
##   (4983 observations deleted due to missingness)
## Multiple R-squared:  0.9878, Adjusted R-squared:  0.9878 
## F-statistic: 6.188e+06 on 1 and 76590 DF, p-value: < 2.2e-16
```

The diagnosis plots can be improved eliminating the values whose cook's distance is above its 0.99775 percentile.

The model is adjusted again.

```
fit.0 <- lm(logW ~ logL, data = data)
summary(fit.0)

##
## Call:
## lm(formula = logW ~ logL, data = data)
##
## Residuals:
##      Min       1Q   Median       3Q      Max 
## -0.34822 -0.03128 -0.00371  0.02813  0.35800 
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)    
## (Intercept) -2.414365   0.001580  -1528   <2e-16 ***
## logL         3.164206   0.001065   2972   <2e-16 ***  
## ---        
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.04723 on 76417 degrees of freedom
## (4983 observations deleted due to missingness)
## Multiple R-squared:  0.9914, Adjusted R-squared:  0.9914 
## F-statistic: 8.834e+06 on 1 and 76417 DF, p-value: < 2.2e-16
```

### Model by year (common intercept, year specific slope)

Estimation of year specific  $b$  (and common  $a$ ). Below the summaries report the model results using two different parametrizations of the same model.

```
fit.YY <- lm(logW ~ logL:Y + logL, data = data)
summary(fit.YY)

##
## Call:
## lm(formula = logW ~ logL:Y + logL, data = data)
##
## Residuals:
##      Min       1Q   Median       3Q      Max 
## -0.34050 -0.03008 -0.00318  0.02734  0.35558 
##
## Coefficients:
##             Estimate Std. Error    t value Pr(>|t|)    
## (Intercept) -2.4146726  0.0016067 -1502.863   < 2e-16 ***
## logL         3.1672825  0.0010830  2924.651   < 2e-16 ***  
## logL:Y.L    -0.0323537  0.0010835   -29.859   < 2e-16 ***  
## logL:Y.Q    -0.0051941  0.0017287    -3.005  0.002660 **  
## logL:Y.C     0.0103104  0.0012312     8.374   < 2e-16 ***  
## logL:Y^4    -0.0075119  0.0016830    -4.463  8.08e-06 *** 
## logL:Y^5     0.0029580  0.0012980     2.279  0.022679 *   
## logL:Y^6    -0.0152811  0.0015889    -9.618   < 2e-16 ***  
## logL:Y^7    -0.0037772  0.0013198    -2.862  0.004212 **  
## logL:Y^8     0.0111245  0.0015192     7.323  2.46e-13 ***
```

```

## logL:Y^9    0.0009670  0.0016171   0.598  0.549845
## logL:Y^10   -0.0012902  0.0013215  -0.976  0.328906
## logL:Y^11    0.0026899  0.0016729   1.608  0.107865
## logL:Y^12   -0.0044380  0.0012643  -3.510  0.000448 ***
## logL:Y^13    0.0050320  0.0017317   2.906  0.003663 **
## logL:Y^14    0.0062401  0.0012337   5.058  4.24e-07 ***
## logL:Y^15   -0.0106238  0.0016891  -6.290  3.20e-10 ***
## logL:Y^16    0.0004268  0.0011648   0.366  0.714058
## logL:Y^17    0.0023624  0.0017300   1.366  0.172081
## logL:Y^18   -0.0119752  0.0012886  -9.293  < 2e-16 ***
## logL:Y^19    0.0103504  0.0015851   6.530  6.62e-11 ***
## logL:Y^20    0.0035704  0.0014325   2.493  0.012686 *
## logL:Y^21   -0.0037021  0.0015072  -2.456  0.014039 *
## logL:Y^22   -0.0048811  0.0016558  -2.948  0.003201 **
## logL:Y^23    0.0096612  0.0015244   6.338  2.34e-10 ***
## logL:Y^24    0.0035673  0.0017844   1.999  0.045598 *
## logL:Y^25   -0.0150287  0.0014054 -10.694  < 2e-16 ***
## logL:Y^26    0.0088529  0.0017529   5.051  4.42e-07 ***
## logL:Y^27   -0.0002690  0.0013994  -0.192  0.847542
## logL:Y^28    0.0090714  0.0020866   4.348  1.38e-05 ***
## logL:Y^29    0.0098937  0.0018067   5.476  4.36e-08 ***
## logL:Y^30    0.0017274  0.0020300   0.851  0.394808
## logL:Y^31   -0.0083472  0.0014994  -5.567  2.60e-08 ***
## logL:Y^32    0.0064917  0.0017542   3.701  0.000215 ***
## logL:Y^33   -0.0053534  0.0020271  -2.641  0.008271 **
## logL:Y^34    0.0066544  0.0030359   2.192  0.028390 *
## logL:Y^35   -0.0033692  0.0027129  -1.242  0.214275
## logL:Y^36   -0.0022713  0.0006514  -3.487  0.000489 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.04599 on 76381 degrees of freedom
##   (4983 observations deleted due to missingness)
## Multiple R-squared:  0.9919, Adjusted R-squared:  0.9919
## F-statistic: 2.519e+05 on 37 and 76381 DF,  p-value: < 2.2e-16

# alternative parameterization
fit.YY.bis <- lm(logW ~ logL:Y, data = data)
summary(fit.YY.bis)

```

```

##
## Call:
## lm(formula = logW ~ logL:Y, data = data)
##
## Residuals:
##      Min       1Q     Median       3Q      Max
## -0.34050 -0.03008 -0.00318  0.02734  0.35558
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)
## (Intercept) -2.414673  0.001607 -1502.9  <2e-16 ***
## logL:Y1982   3.164556  0.001637  1932.9  <2e-16 ***
## logL:Y1983   3.170155  0.001926  1645.6  <2e-16 ***
## logL:Y1984   3.171420  0.001356  2339.2  <2e-16 ***
## logL:Y1985   3.179782  0.001216  2615.3  <2e-16 ***

```

```

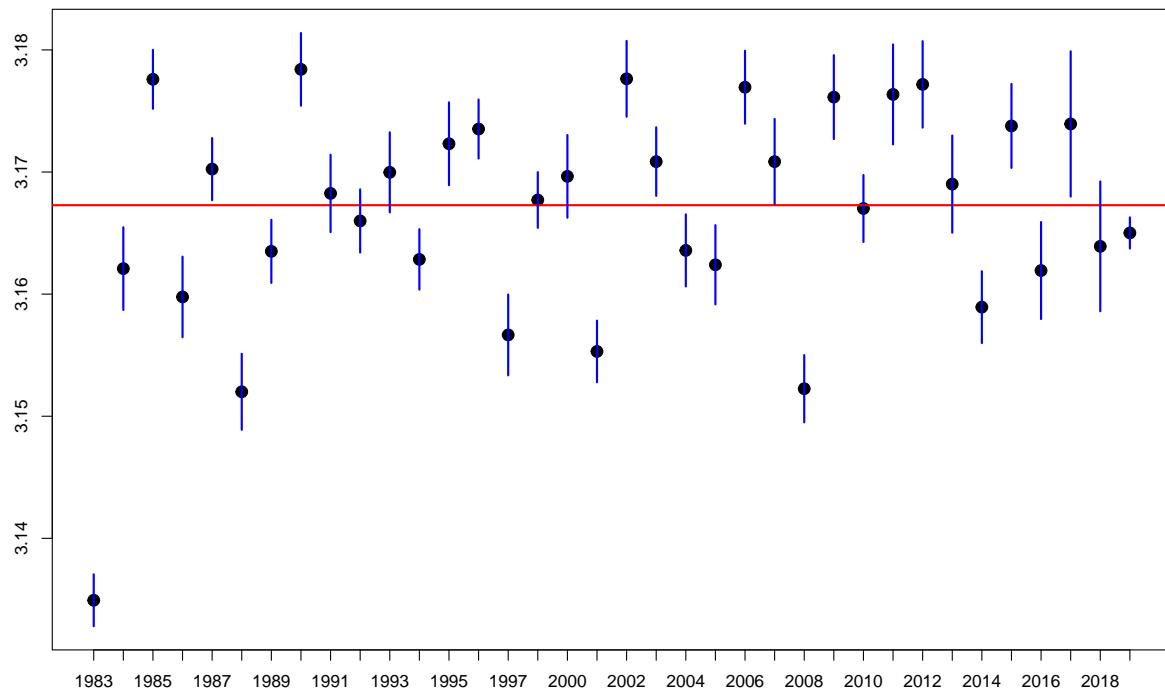
## logL:Y1986 3.173834 0.001631 1945.4 <2e-16 ***
## logL:Y1987 3.172121 0.002379 1333.1 <2e-16 ***
## logL:Y1988 3.195030 0.001454 2196.7 <2e-16 ***
## logL:Y1989 3.174024 0.001374 2309.6 <2e-16 ***
## logL:Y1990 3.178302 0.001579 2013.4 <2e-16 ***
## logL:Y1991 3.168156 0.001634 1939.2 <2e-16 ***
## logL:Y1992 3.157475 0.002647 1192.6 <2e-16 ***
## logL:Y1993 3.168621 0.002456 1290.1 <2e-16 ***
## logL:Y1994 3.175160 0.001532 2072.9 <2e-16 ***
## logL:Y1995 3.176700 0.001633 1945.2 <2e-16 ***
## logL:Y1996 3.171229 0.001387 2286.6 <2e-16 ***
## logL:Y1997 3.160179 0.001728 1828.8 <2e-16 ***
## logL:Y1998 3.161178 0.006667 474.2 <2e-16 ***
## logL:Y2000 3.187652 0.004069 783.3 <2e-16 ***
## logL:Y2001 3.170237 0.001847 1716.6 <2e-16 ***
## logL:Y2002 3.163135 0.001715 1843.9 <2e-16 ***
## logL:Y2003 3.169709 0.001350 2347.2 <2e-16 ***
## logL:Y2004 3.174690 0.001206 2632.7 <2e-16 ***
## logL:Y2005 3.165045 0.001168 2709.8 <2e-16 ***
## logL:Y2006 3.161249 0.001201 2632.0 <2e-16 ***
## logL:Y2007 3.161313 0.001151 2747.8 <2e-16 ***
## logL:Y2008 3.153412 0.001166 2704.9 <2e-16 ***
## logL:Y2009 3.158032 0.001185 2664.4 <2e-16 ***
## logL:Y2010 3.159046 0.001174 2691.2 <2e-16 ***
## logL:Y2011 3.161997 0.001180 2679.9 <2e-16 ***
## logL:Y2012 3.161874 0.001235 2560.2 <2e-16 ***
## logL:Y2013 3.159695 0.001182 2672.9 <2e-16 ***
## logL:Y2014 3.161606 0.001186 2665.1 <2e-16 ***
## logL:Y2015 3.166102 0.001211 2614.3 <2e-16 ***
## logL:Y2016 3.164287 0.001199 2638.4 <2e-16 ***
## logL:Y2017 3.159834 0.001214 2603.3 <2e-16 ***
## logL:Y2018 3.157208 0.001267 2490.9 <2e-16 ***
## logL:Y2019 3.155406 0.001340 2354.3 <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.04599 on 76381 degrees of freedom
##   (4983 observations deleted due to missingness)
## Multiple R-squared: 0.9919, Adjusted R-squared: 0.9919
## F-statistic: 2.519e+05 on 37 and 76381 DF, p-value: < 2.2e-16

```

### Plots (model by year)

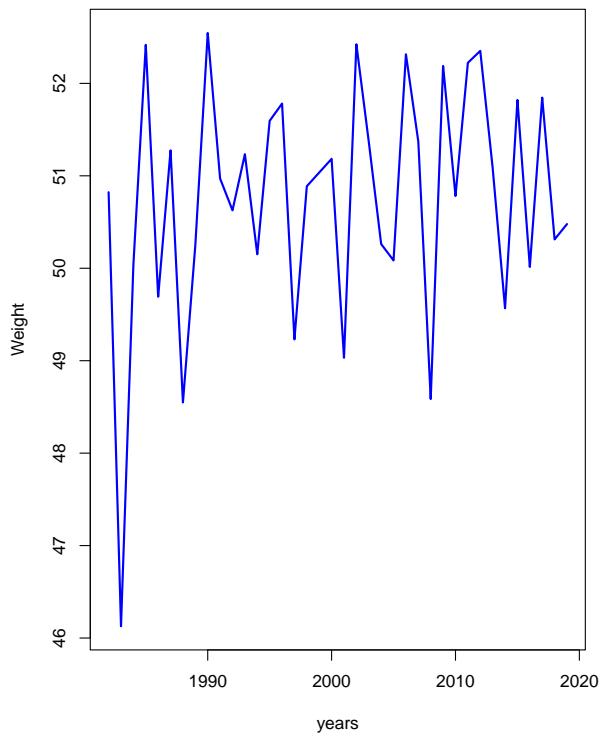
Comparison of the  $b$  year specific estimates (black point) respect to the  $b$  estimate at the first year of our time series (1982, red horizontal line). The 95% confidence intervals are plotted using a vertical line.

Slope estimates (Reference Value 1982, horizontal line)

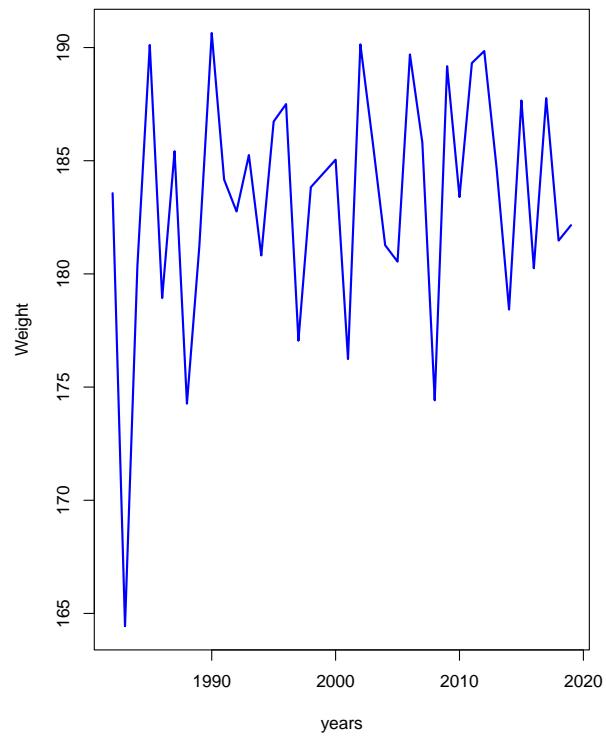


Predicted weights along the years for the length values of the sequence 20, 30, ..., 100 cm. The plots show that the predicted weight series are more or less stable for the last years (and for all the lengths). Then, it seems that year specific  $b$  estimates are not required.

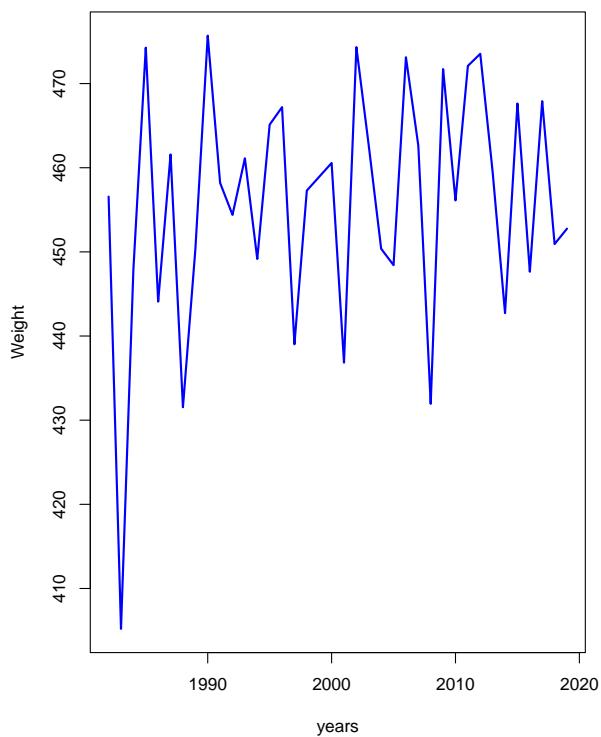
**Prediction of weight for 20 cm through years**



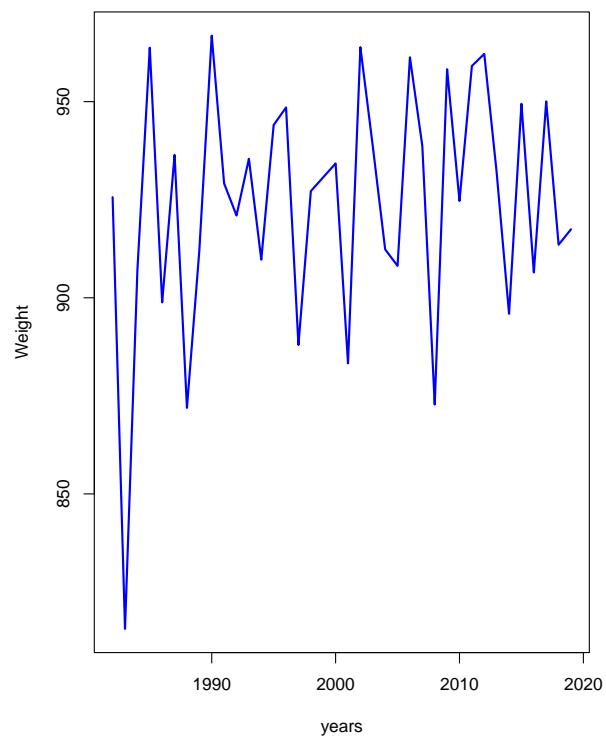
**Prediction of weight for 30 cm through years**



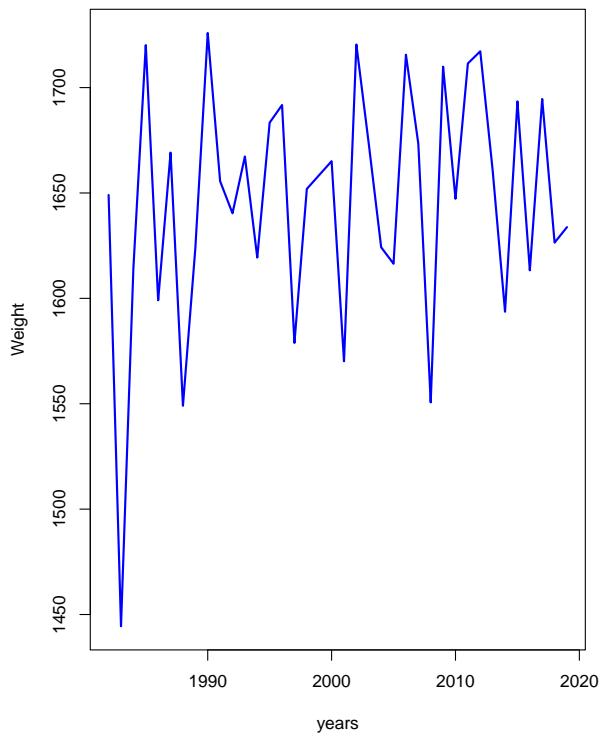
**Prediction of weight for 40 cm through years**



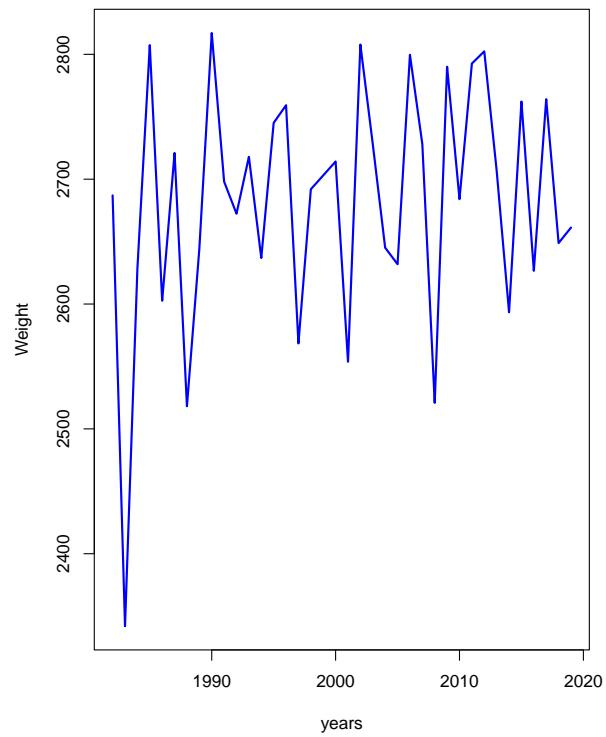
**Prediction of weight for 50 cm through years**



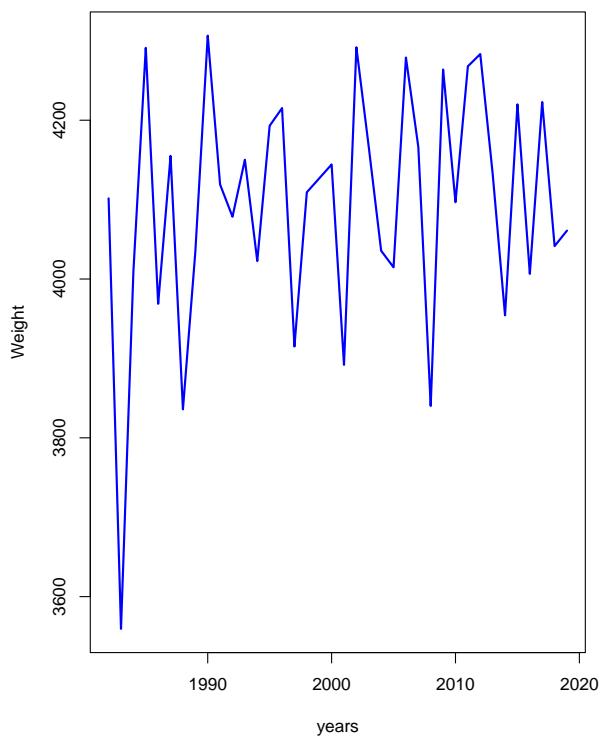
**Prediction of weight for 60 cm through years**



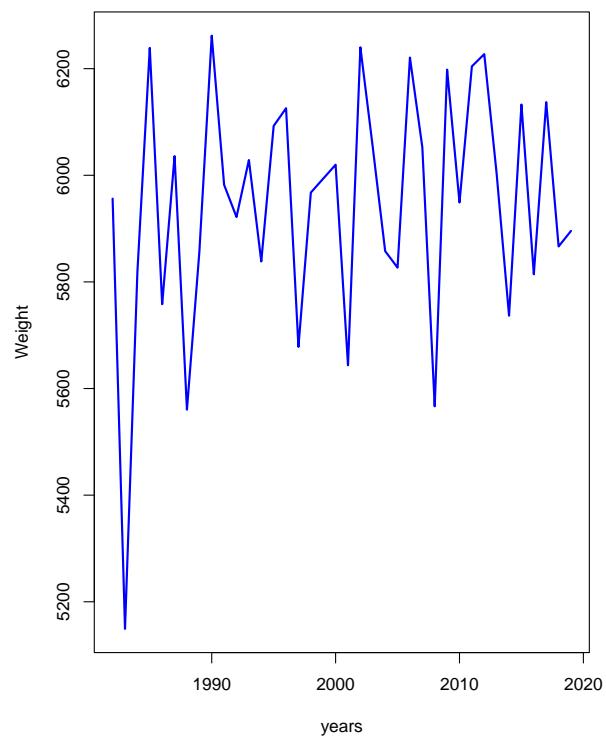
**Prediction of weight for 70 cm through years**



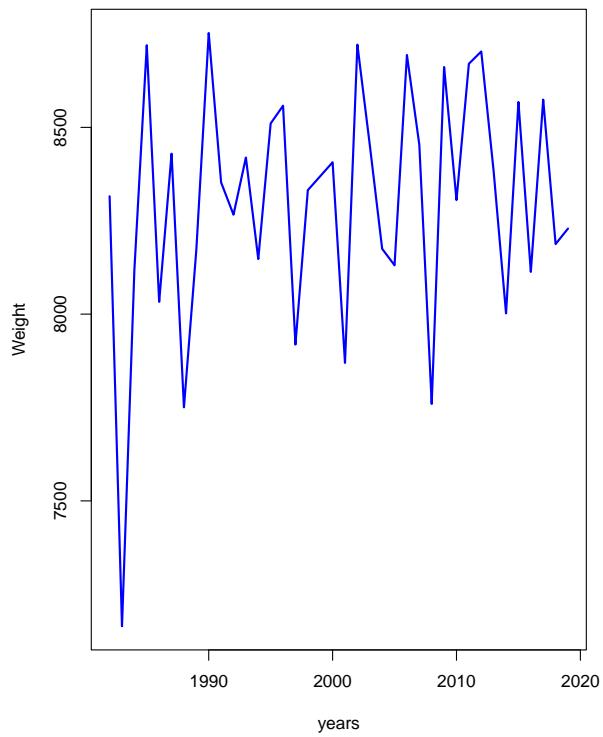
**Prediction of weight for 80 cm through years**



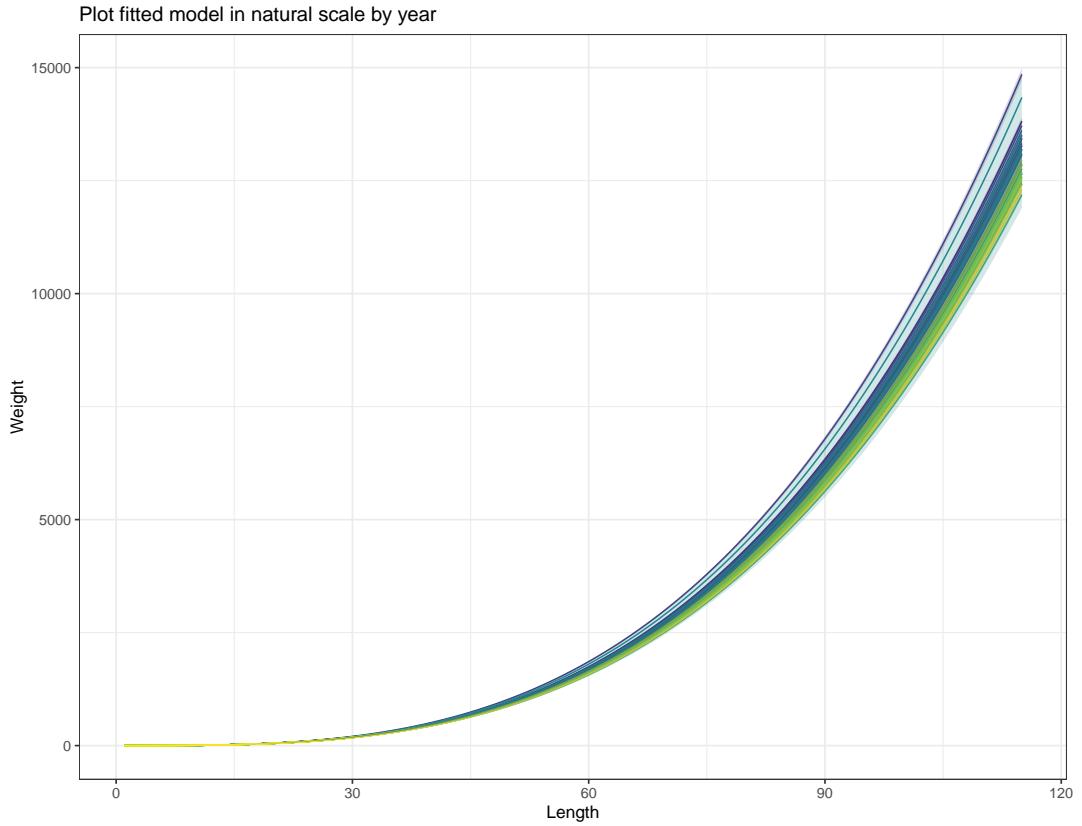
**Prediction of weight for 90 cm through years**



### Prediction of weight for 100 cm through years



Additionally to previous plots, the next plot show the weight-length relation curves for each year.



## Model by sex

Estimation of global  $a$  and  $b$  by sex. Below the summaries report the model results using two different parametrizations of the same model.

```
fit.SSS <- lm(logW ~ logL*sex, data = data)
summary(fit.SSS)

##
## Call:
## lm(formula = logW ~ logL * sex, data = data)
##
## Residuals:
##      Min       1Q   Median       3Q      Max 
## -0.33954 -0.02925 -0.00326  0.02611  0.35808 
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)    
## (Intercept) -2.438396  0.002847 -856.60 <2e-16 ***
## logL         3.180831  0.001844 1725.16 <2e-16 ***
## sexI          0.153870  0.007972   19.30 <2e-16 ***
## sexM          0.054033  0.004892   11.04 <2e-16 ***
## logL:sexI   -0.123814  0.006235  -19.86 <2e-16 ***
## logL:sexM   -0.043534  0.003263  -13.34 <2e-16 ***
## ---        
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.0452 on 62656 degrees of freedom
```

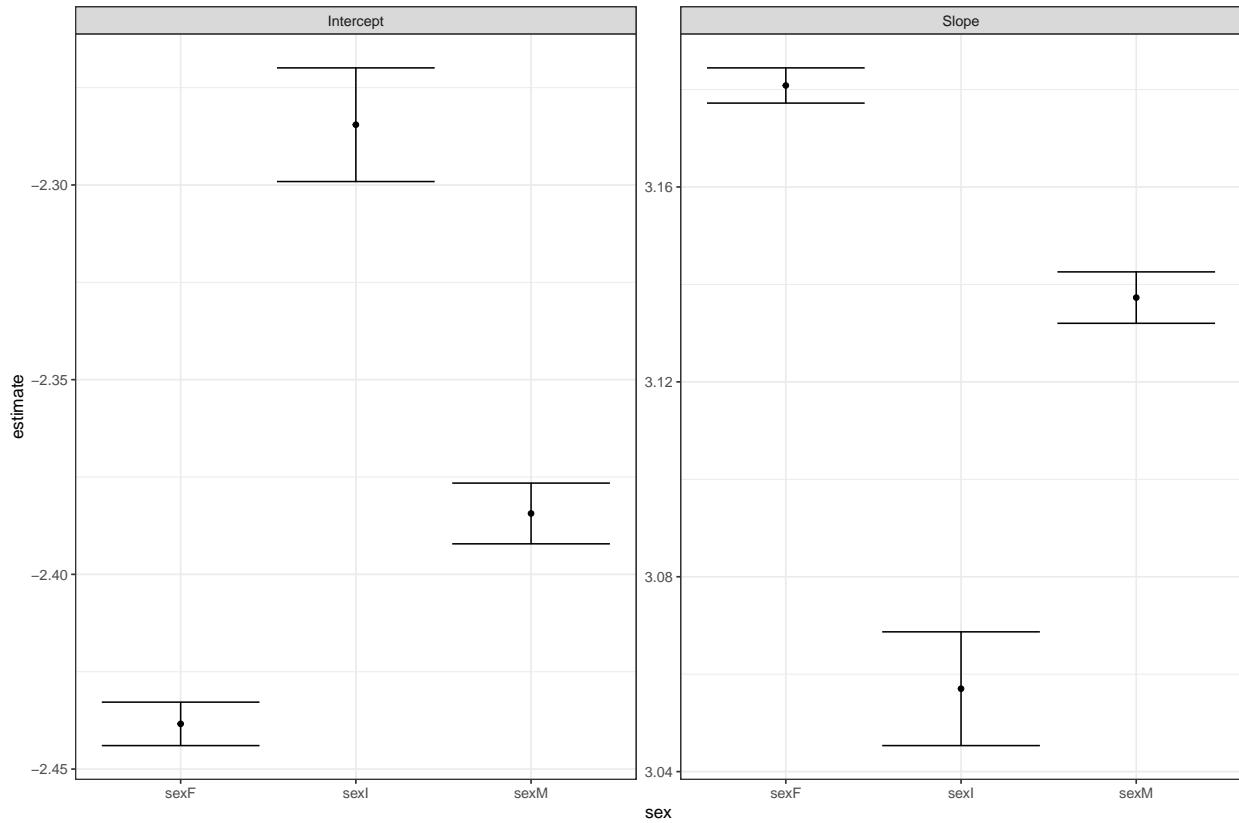
```

##   (3153 observations deleted due to missingness)
## Multiple R-squared:  0.9921, Adjusted R-squared:  0.9921
## F-statistic: 1.577e+06 on 5 and 62656 DF, p-value: < 2.2e-16
fit.SSS.bis <- lm(logW ~ logL*sex -1 -logL, data = data)
summary(fit.SSS.bis)

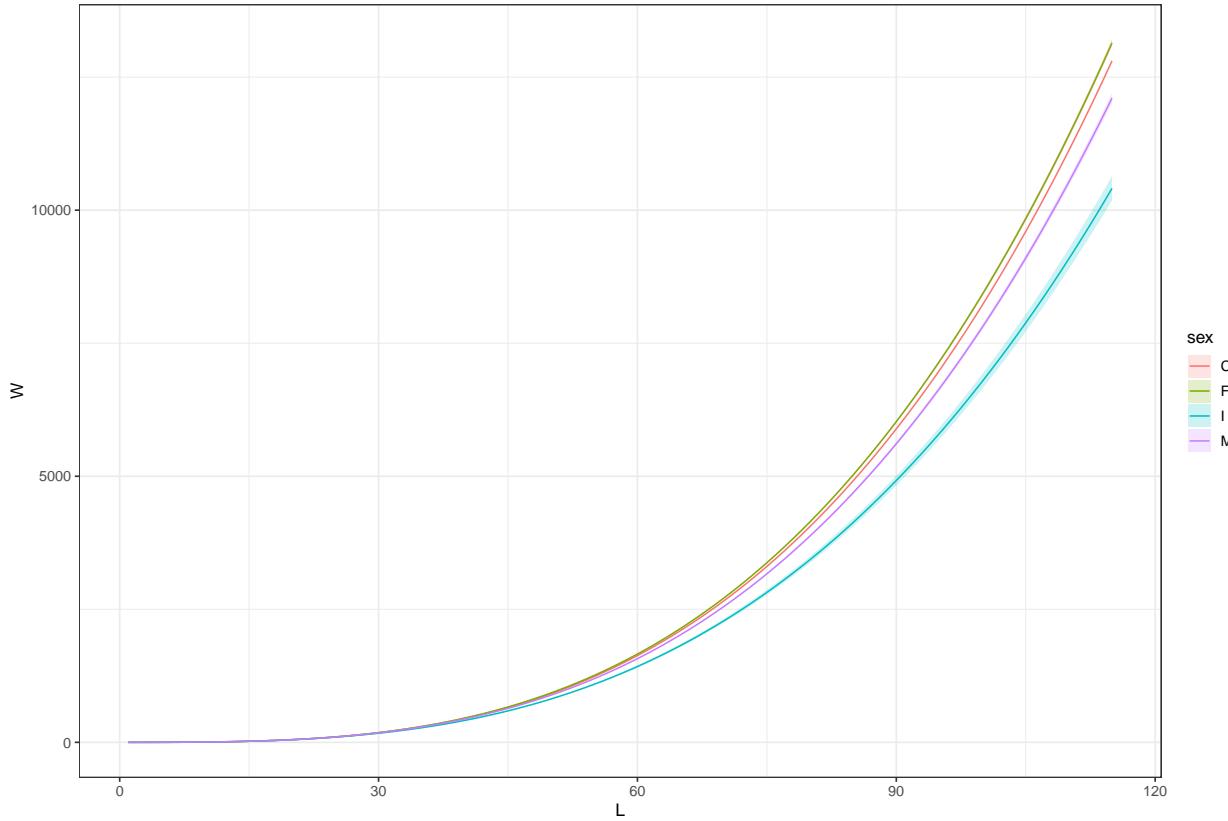
##
## Call:
## lm(formula = logW ~ logL * sex - 1 - logL, data = data)
##
## Residuals:
##       Min     1Q Median     3Q    Max 
## -0.33954 -0.02925 -0.00326  0.02611  0.35808 
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)    
## sexF      -2.438396  0.002847 -856.6   <2e-16 ***
## sexI      -2.284526  0.007446 -306.8   <2e-16 *** 
## sexM      -2.384363  0.003979 -599.3   <2e-16 *** 
## logL:sexF  3.180831  0.001844 1725.2   <2e-16 *** 
## logL:sexI  3.057018  0.005956  513.3   <2e-16 *** 
## logL:sexM  3.137297  0.002693 1165.1   <2e-16 *** 
## ---      
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 
##
## Residual standard error: 0.0452 on 62656 degrees of freedom
##   (3153 observations deleted due to missingness)
## Multiple R-squared:  0.9996, Adjusted R-squared:  0.9996 
## F-statistic: 2.634e+07 on 6 and 62656 DF, p-value: < 2.2e-16

```

Plot of the  $a$  and  $b$  estimates for each sex (F=females, M=males, I=inmatures) with the corresponding confidence interval.



Plot of the sex specific length-weight curves (C means global model, that is, no sex differences). The results show that the females and males curves are almost the same for the range of lengths for which data of both sexes is available, hence it lead to conclude that the sex differences are not relevant enough to propose sex-specific estimates.



## Final proposal

We propose to update the  $a$  and  $b$  values using the estimates derived from the global model using data from 2003 to 2019. The selection of years has been carried out focusing on the sampling sizes of both countries.

1999 estimates:

$a = 0.00659$   $b = 3.01721$

Updated:

$a=0.00377$   $b=3.16826$

```
data=subset(data,data$year>2002)
```

```
fit.0 <- lm(logW ~ logL, data = data)
```

```
summary(fit.0)
```

```
##  
## Call:  
## lm(formula = logW ~ logL, data = data)  
##  
## Residuals:  
##      Min       1Q   Median       3Q      Max  
## -0.34459 -0.02976 -0.00350  0.02659  0.36163  
##  
## Coefficients:  
##             Estimate Std. Error t value Pr(>|t|)  
## (Intercept) -2.423832   0.001678  -1445   <2e-16 ***  
## logL         3.168263   0.001140    2780   <2e-16 ***  
## ---
```

```

## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.04565 on 62660 degrees of freedom
##   (3153 observations deleted due to missingness)
## Multiple R-squared:  0.992, Adjusted R-squared:  0.992
## F-statistic: 7.727e+06 on 1 and 62660 DF, p-value: < 2.2e-16

```

Comparison of the updated weight-length curve with the 1999 curve. The confidence interval is plotted in green. The grey points correspond to the observed data. The curves differs for large length values for which almost no data is available.

